

A comparison of spectral estimation methods for the analysis of chaotic and stochastic dynamical systems

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What is power spectrum?

Conceptual



- (For ease of exposition) Start with a continuous time stochastic process $X(t)$

- We have

- ▶ $\mu_t = \mathbb{E}X(t)$

- ▶ $C_X(t, s) = \mathbb{E}(X(t) - \mu_t)(X(s) - \mu_s)^*$

- The process is wide-sense stationary if

$$\mu_t = \mu \quad (\text{constant})$$

and

$$C_X(t, s) = C_X(t - s) \quad (\text{depends on on lag})$$

- Center $X(t)$

$$X(t) \leftarrow X(t) - \mu$$

What is power spectrum?

Conceptual



- From signals and systems we get the terms

- ▶ energy

$$\text{Total energy of } X(t) \text{ over } (t_1, t_2) = \int_{t_1}^{t_2} |X(t)|^2 dt$$

- ▶ power

$$\text{Total power of } X(t) \text{ over } (t_1, t_2) = \frac{1}{t_2 - t_1} \int_{t_1}^{t_2} |X(t)|^2 dt$$

- If $X(t)$ is **deterministic** and **periodic** with period $2T$

- ▶ $X(t) = \sum_{n=0}^{\infty} c_n e^{j\pi n t / T}$

- ▶ So, total power over $(-T, T) = \frac{1}{2T} \int_{-T}^T |X(t)|^2 dt = \sum_{n=0}^{\infty} |c_n|^2$

What is power spectrum?

Conceptual



- Example:

- ▶ If $X(t) = c_n e^{i\pi n t/T}$ then total power = $|c_n|^2$

interpretation

$|c_n|^2$ = contribution to the total power from the term in the Fourier series of $X(T)$ with frequency $n/2T$ Hz (or angular frequency of $\pi n/T$ radians per second).

What is power spectrum?

Conceptual



- If $X(t)$ is **deterministic** and **nonperiodic**

- ▶ $X(t) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} G(\omega) e^{i\omega t} d\omega \quad (X \in L^2(\mathbb{R}))$ Fourier integral

- ▶ So, total energy over $\mathbb{R} = \int_{-T}^T |X(t)|^2 dt = \int_{-T}^T |G(\omega)|^2 d\omega$

interpretation

$|G(\omega)|^2 d\omega =$ contribution to the total energy from components of $X(t)$ whose frequencies lie between ω and $\omega + d\omega$ radians per second.

What is power spectrum?

Conceptual



- If $X(t)$ is **stochastic** and **stationary**

- ▶ take a realization of $X(t)$ $X \notin L^2$

- ▶ $X_T(t) = X(t)|_{[-T, T]}$ $X_T \in L^2$

- ▶ $X_T(t) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} G_T(\omega) e^{i\omega t} d\omega$ where $G_T(t) = \frac{1}{\sqrt{2\pi}} \int_{-T}^T X(\omega) e^{-i\omega t} d\omega$

- ▶ So, we have an interpretation of $|G_T(\omega)|^2 d\omega$

interpretation

$|G_T(\omega)|^2 d\omega =$ contribution to the total energy from components of $X_T(t)$ whose frequencies lie between ω and $\omega + d\omega$ radians per second.

What is power spectrum?

Conceptual



Program in Applied
Mathematics

interpretation

$\lim_{T \rightarrow \infty} \frac{|G_T(\omega)|^2}{2T} d\omega =$ contribution to the total power from components of $X_T(t)$ whose frequencies lie between ω and $\omega + d\omega$ radians per second.

$$S_X(\omega) = \lim_{T \rightarrow \infty} \mathbb{E} \frac{|G_T(\omega)|^2}{2T}$$

interpretation

$S_X(\omega)d\omega =$ average (over all realizations) of the contribution to the total power from components in $X(t)$ with frequencies between ω and $\omega + d\omega$ radians per second.

What is power spectrum?

Operational



Program in Applied
Mathematics

- Start with X_n
 - ▶ a discrete-time stochastic process,
 - ▶ wide-sense stationary, and
 - ▶ centered.
- The power spectrum $S_X(\omega)$ is define by

$$S_X(\omega) = \sum_{n=-\infty}^{\infty} C_X(n)e^{-i\omega n} = \mathcal{F}\{C_X\}(\omega) = \widehat{C_X}(\omega)$$

where $C_X(n) = \mathbb{E}X_n X_0^*$ (Fourier transform of the autocovariance function)

- The z-spectrum $\bar{S}_X(z)$ is define by

$$\bar{S}_X(\omega) = \sum_{n=-\infty}^{\infty} C_X(n)z^{-n} = \mathcal{Z}\{C_X\}(\omega)$$

(z-transform of the autocovariance function)

What is power spectrum?

Observation



- Observe that, by the inverse Fourier transform formula

$$\text{var}(X) = C_X(0) = \frac{1}{2\pi} \int_{-\pi}^{\pi} S_X(\omega) e^{i\omega 0} d\omega = \frac{1}{2\pi} \int_{-\pi}^{\pi} S_X(\omega) d\omega$$

So that spectrum given the distribution of variance among the frequencies.

We are given data, x_n , for $n = 1, 2, 3, \dots, N$

- Assume it is be a realization of the discrete-time process X_n or observations of a continuous time process X_{t_n} .
- Assume the process X_n is stationary

How do we estimate μ ?

By virtue of stationary

$$\mu = \mathbb{E}X_n \approx \frac{1}{N} \sum_{n=1}^N x_n =: \tilde{\mu}$$

How do we estimate $R_X(n)$?

Again, by virtue of stationary

$$R_X(n) = \mathbb{E}[(X_n - \mu)(X_0 - \mu)^*] \approx \frac{1}{N} \sum_{j=1}^{N-n} (x_{n+j} - \tilde{\mu})(x_j - \tilde{\mu})^* =: \tilde{R}_X(n)$$

Spectrum Estimation (sample spectrum)

Periodogram



How do we estimate $S_X(\omega)$? (assume X_n is mean zero)

Periodogram: (direct approach)

$$\begin{aligned}\tilde{S}_X(\omega) &= \sum_n \tilde{R}_X(n) e^{-in\omega} \\ &= \sum_n \frac{1}{N} \sum_j x_{n+j} x_j^* e^{-in\omega} \\ &= \frac{1}{N} \sum_k \sum_j x_k x_j^* e^{-ik\omega} e^{ij\omega} \\ &= \frac{1}{N} \left(\sum_k x_k e^{-ik\omega} \right) \left(\sum_j x_j e^{-ij\omega} \right)^* \\ &= \frac{1}{N} \hat{x}(\omega) \hat{x}(\omega)^* = \frac{1}{N} |\hat{x}(\omega)|^2 \quad (= \text{abs2} . (\text{fft}(\mathbf{x})) / N)\end{aligned}$$

Asymptotically unbiased but inconsistent (the variance does not vanish as N gets large).

How do else we estimate $S_X(\omega)$?

Bartlett's smoothing procedure: cut up the timeseries into k blocks. And approximate the peridogram $\tilde{S}_X^{(j)}(\omega)$ for each block of data $j = 1, 2, \dots, k$.

$$\tilde{S}_X(\omega) = \frac{1}{k} \sum_{J=1}^k \tilde{S}_X^{(j)}(\omega)$$

This procedure allows us to control the variance, but at the expense of bias. This procedure can be generalized.

General class of smoothed spectral estimators:

Bartlett:

$$\tilde{S}_X(\omega) = \frac{1}{k} \sum_{n=-k}^k \left(1 - \frac{|n|}{k}\right) \tilde{R}_X(n) e^{-in\omega}$$

General:

$$\tilde{S}_X(\omega) = \frac{1}{k} \sum_{n=-\infty}^{\infty} w(n) \tilde{R}_X(n)$$

with

- (1) $w(0) = 1$
- (2) $w(n) = w(-n)$
- (3) $w(n) = 0, \quad |n| \geq k, \quad k < N$

w is called a windowing function.

Most common window functions,

Bartlett:

$$w(n) = \begin{cases} 1 - \frac{|n|}{k}, & |n| \leq k \\ 0, & |n| > k \end{cases}$$

Tukey:

$$w(n) = \begin{cases} \frac{1}{2} \left(1 + \cos \frac{\pi n}{k} \right), & |n| \leq k \\ 0, & |n| > k \end{cases}$$

Parzen:

$$w(n) = \begin{cases} 1 - 6 \left(\frac{n}{k} \right)^2 + 6 \left(\frac{|n|}{k} \right)^3, & |n| \leq k/2 \\ 2 \left(1 - \frac{|n|}{k} \right)^3, & k/2 < |n| \leq k \\ 0, & |n| > k \end{cases}$$

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What's Welch?

- Lattice of autoregressive model of order p , $p = 1, \dots, p_{\max}$.

$$X_n + a_{1,p}X_{n-1} + \dots + a_{p,p}X_{n-p} = e_{n,p}$$

$$a_{p,p}X_n + a_{p-1,p}X_{n-1} + \dots + X_{n-p} = r_{n,p}$$

- Minimize $E_p = \sum_{n=p+1}^N (|e_{n,p}|^2 + |r_{n,p}|^2)$

- Because of these are linear least squares coefficients

$$e_{n,p} = e_{n,p-1} + a_{p,p}r_{n-1,p-1}$$

$$r_{n,p} = a_{p,p}e_{n,p-1} + r_{n-1,p-1}$$

$$\begin{pmatrix} a_{1,p} \\ \vdots \\ a_{p-1,p} \end{pmatrix} = \begin{pmatrix} a_{1,p-1} \\ \vdots \\ a_{p-1,p-1} \end{pmatrix} + a_{p,p} \begin{pmatrix} a_{p-1,p-1} \\ \vdots \\ a_{1,p-1} \end{pmatrix}$$

- We end up with

$$a_{p,p} = -\frac{2 \sum_{n=p+1}^N e_{n,p-1} r_{n-1,r-1}}{\sum_{n=p+1}^N (e_{n,p-1}^2 + r_{n-1,r-1}^2)}$$

- For a given p the spectral estimate

$$\hat{S}_X^{\text{Burg}} = \frac{\sigma_p^2}{|A(\omega)|^2} \quad \text{where} \quad A(\omega) = \sum_{k=0}^p a_{k,p} e^{-ik\omega}$$

there are a number of information criteria that can be used to select the order p .

Example 1: AR(2) Signal



Let us consider the stationary autoregressive process of order 2, with poles at $r_1 = .5, r_2 = -.8$

$$Y_n = (r_1 + r_2)Y_{n-1} - r_1r_2Y_{n-2} + e_n = -0.3Y_{n-1} + 0.4Y_{n-2} + e_n, \quad \text{for } n > -\infty$$

for e_n are i.i.d. standard normal random variables.

One way to compute the z-spectrum is as follows. Recognize,

$$(r \star Y)_n = Y_n + 0.3Y_{n-1} - 0.4Y_{n-2} = e_n,$$

$$r = (\dots, 0, \boxed{1}, 0.3, -0.4, 0, \dots)$$

So that,

$$1 = S_e(z) = S_{(r \star Y)} = \bar{r}(z)S_Y(z)\bar{r}^*(z^{-*})$$

and

$$S_Y(z) = \frac{1}{\bar{r}(z)\bar{r}^*(z^{-*})} = \frac{1}{(1 - 0.5z^{-1})(1 + 0.8z^{-1})(1 - 0.5z)(1 + 0.8z)}$$

Example 1: AR(2) Signal

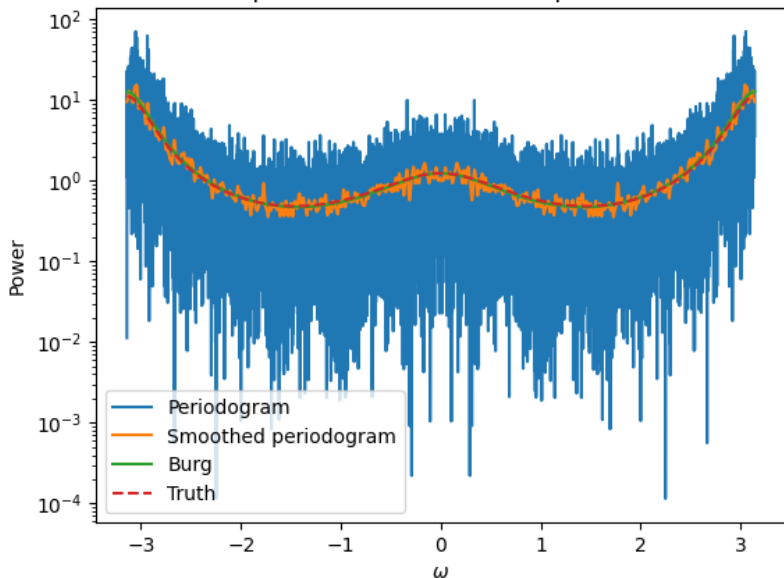


```
1 using DSP, PyPlot, FFTW
2 at = include("../Tools/AnalysisToolbox.jl")
3 se = include("../Tools/SpecEst.jl")
4 bg = include("../Tools/Burg.jl")
5
6 r1 = .5; r2 = -.8
7 r = [1, -(r1 + r2), r1*r2]
8 f(z) = sum(r[j]*z^(1-j) for j=1:3)
9
10 N = 10^4
11 fil = ZeroPoleGain(zeros(0),[r1,r2],1)
12 y = filt(fil,randn(N))
13
14 Sy_per = abs2.(fft(y))/N
15 Sy_num_gb = se.spec_GB(at.rowmatrix(y); Nex = N).S[:];
16 Sy_num_burg = bg.spec_mesa_sc(at.rowmatrix(y); Nex = N, p_max = 100).S[:];
17 Sy_ana = map(z -> 1/abs2(f(z)),exp.(2pi*im*(0:N-1)/N))
18
19 0 = 2pi*(0:N-1)/N .- pi
20 title("Spectral estimates of AR(2) process")
21 semilogy(0,ifftshift(Sy_per), label = "Periodogram")
22 semilogy(0,ifftshift(Sy_num_gb), label = "Smoothed periodogram")
23 semilogy(0,ifftshift(Sy_num_burg), label = "Burg")
24 semilogy(0,ifftshift(Sy_ana), "--", label = "Truth")
25 ylabel("Power")
26 xlabel(L"\omega")
27 legend()
```

Example 1: AR(2) Signal



Spectral estimates of AR(2) process



Example 2: AR(10) Signal

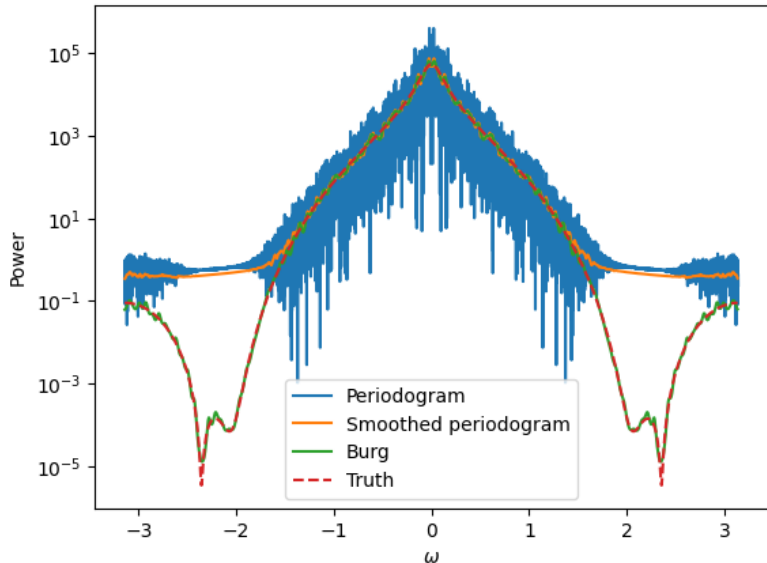


```
1 using DSP, PyPlot, FFTW, Polynomials
2 at = include("../Tools/AnalysisToolbox.jl")
3
4 Poles = [.9];
5
6 Zeros = [.99exp(1im*3pi/4); .99exp(-1im*3pi/4);
7         .9exp(1im*3pi/2*.9); .9exp(-1im*3pi/2*.9);
8         .9exp(1im*3pi/2*.9); .9exp(-1im*3pi/2*.9)]
9
10 spec = x -> at.poles2spec(Poles)(x) * at.zeros2spec(Zeros)(x)
11 spec = spec * at.expi
12
13 N = 10^4
14 X = at.ARMA_gen(; steps = N, Poles, Zeros, r1 = true)
15 X = at.rowmatrix(X)
16
17 Nex = 1000
18 L = 500
19
20 fgrid = se.0(Nex) .- pi
21 SX_ana = spec.(fgrid);
22
23 SX_per = abs2.(fft(X[:]))/N
24 SX_num_gb = se.spec_GB(X; L, Nex).S[:];
25 SX_num_burg = bg.spec_mesa_sc(X; Nex, p_max = 100).S[:];
26
27 0 = 2pi*(0:N-1)/N .- pi
28 title("Spectral estimates of ARMA(1,6) process, N = $N")
29 semilogy(0,ifftshift(SX_per), label = "Periodogram")
30 semilogy(fgrid,ifftshift(SX_num_gb), label = "Smoothed periodogram")
31 semilogy(fgrid,ifftshift(SX_num_burg), label = "Burg")
32 semilogy(fgrid,SX_ana, "--", label = "Truth")
33 ylabel("Power")
34 xlabel(L"\omega")
35 legend()
```

Example 2: AR(10) Signal



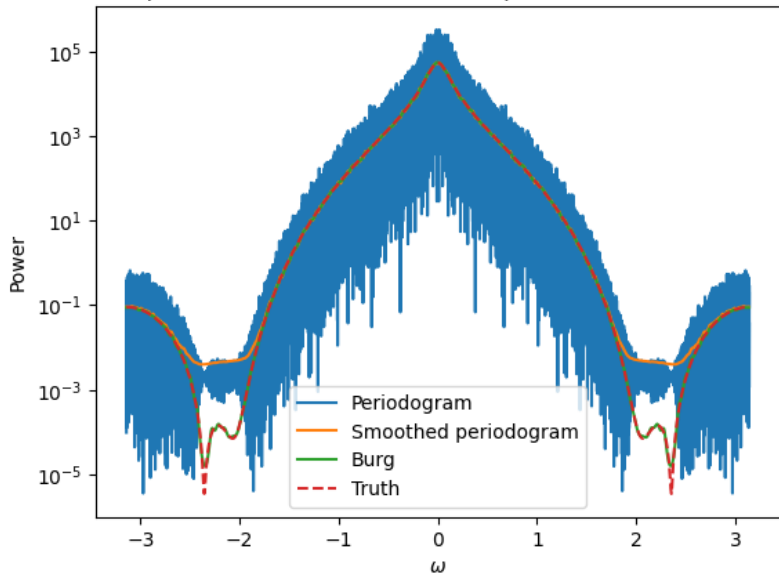
Spectral estimates of ARMA(1,6) process, $N = 10000$



Example 2: AR(10) Signal



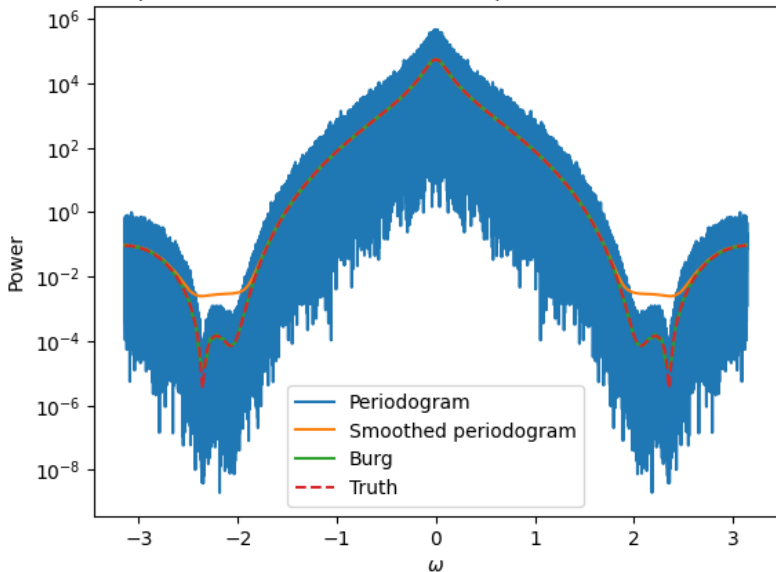
Spectral estimates of ARMA(1,6) process, $N = 100000$



Example 2: AR(10) Signal



Spectral estimates of ARMA(1,6) process, $N = 1000000$



Now, spectral factorization!

- For the numerical spectral factorization we do, assume $S_X(\omega)$ is rational.
- Because is a power spectrum, $S_X(\omega) \geq 0$ on $-\pi$ to π .
- So, we can factor $S_X(\omega) = L(\omega)L^*(\omega)$
 - ▶ $L(\omega) = \sqrt{S_X(\omega)}$
 - ▶ $L(\omega)$ is minimum phase
- minimum phase:
 - ▶ $\bar{S}_X(z) = \bar{L}(z)\bar{L}^*(z^{-*})$
 - ▶ $L(\omega) = \bar{L}(e^{j\omega})$
 - ▶ (minimum phase) $\bar{L}(z)$ and $\bar{L}^{-1}(z)$ are analytic on and outside the unit circle. ($\bar{L}(z)$ has all it's poles strictly inside the unit circle)



- Write $L^{-1}(\omega) = \sum_{n=-\infty}^{\infty} w_n e^{-i\omega n}$
- w_n is the Fourier coefficients of $L^{-1}(\omega)$

- It can be shown that

$$S_{w*X}(\omega) = L^{-1}(\omega)S_X(\omega)L^{-*}(\omega) = S_X(\omega)/S_X(\omega) = 1$$

- w is a whitening filter for X .



- If $L(\omega)$ is minimum phase, so is $L^{-1}(\omega)$ and

$$L^{-1}(\omega) = \sum_{n=0}^{\infty} w_n e^{-i\omega n}$$

So, $w_n = 0$ for $n < 0$, we say w is causal.

Most Numerical algorithms assume $S(z)$ is rational and has the form of a Laurent Polynomial meaning it may be written as

$$S(z) = \sum_{n=-m}^m c_n z^{-n} \quad \text{with } c_n = c_{-n}^*.$$

If this is assumed it may be shown that

$$S^+(z) = \sum_{n=1}^m L_n z^n \quad \text{and} \quad S^-(z) = \sum_{n=1}^m L_n^* z^{-n}$$

(this is what we assume here) Algorithms that use Toeplitz matrices.

- Bauer
- Schur
- Levinson-Durbin

Algorithms that use state-space formulations.

- Riccati Equation
- Kalman Filter
- Chadrachar-Kailath-Morf-Sidhu (CKMS)

For the DFT which is used frequently in this work I use `fft` from `FFTW.jl` which is a Julia wrapper for the `FFTW` library written in C.

Here is what it does:

$$v_k = \text{fft}(u)_k = \sum_{j=1}^N u_j e^{-\frac{2\pi i}{N}(j-1)(k-1)}$$
$$u_j = \text{ifft}(v)_j = \frac{1}{N} \sum_{k=1}^N v_k e^{\frac{2\pi i}{N}(k-1)(j-1)}$$

Here is why I use it so much:

Suppose we have the function $S(z) = \sum_{j=1}^N c_j z^{-(j-1)}$ which we wish to evaluate at N equally-spaced, unit-circle points $z_k = e^{\frac{2\pi i}{N}(k-1)}$ for $k = 1, \dots, N$. We need only use `fft` to get

$$S(z_k) = \sum_{j=1}^N c_j e^{-\frac{2\pi i}{N}(j-1)(k-1)} = \text{fft}(c)_k.$$

So, given a causal finite impulse response (FIR) filter ℓ , it's transfer function $L(z)$ evaluated at N_{ex} evenly distributed points on the unit circle is the array

$$\left(L(z) : z = e^{2\pi ik/N_{ex}} \text{ for } k = 0, \dots, N_{ex} - 1 \right) = \text{fft}([\ell; \text{zeros}(N_{ex} - \text{length}(\ell))])$$

The first entry corresponds to $L(1)$ and the points go counterclockwise. So, to get an approximate inverse of an causal FIR a filter.

- Parzen (1957): Error in Bartlett mainly due to variance.
- Control Variates
 - ▶ Estimate an expectation $\mu = \mathbb{E}X$, of some random variable X
 - ▶ Take n IID samples X_i of X

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i.$$

- ▶ $\hat{\mu}$ is an unbiased estimator of μ
- ▶ $\text{var}(\hat{\mu}) = \text{var}(X)/n$.

- Suppose Y (mean zero) correlated with X .
- Take n IID samples Y_i of Y
- Consider

$$\hat{\mu}^{\text{cv}} = \frac{1}{n} \sum_{i=1}^n X_i - \alpha Y_i$$

- ▶ $\hat{\mu}^{\text{cv}}$ is an unbiased estimator of μ
- ▶ $\text{var}(\hat{\mu}^{\text{cv}}) = \frac{1}{n} \text{var}(X - \alpha Y)$ and

$$\begin{aligned} \text{var}(X - \alpha Y) &= \mathbb{E}(X - \alpha Y - \mu)(X - \alpha Y - \mu)^* \\ &= \mathbb{E}(X - \mu)(X - \mu)^* - \alpha \mathbb{E}Y(X - \mu)^* - \mathbb{E}(X - \mu)Y^* \alpha^* + \alpha \mathbb{E}YY^* \\ &= \text{var}(X) - 2\mathcal{R}\{\alpha \text{cov}(Y, X)\} + |\alpha|^2 \text{var}(Y) \end{aligned}$$

- ▶ minimizer $\alpha = \frac{\text{cov}(Y, X)^*}{\text{var}(Y)} = \frac{\text{cov}(X, Y)}{\text{var}(Y)}$

- So, for this $\alpha = \frac{\text{cov}(X, Y)}{\text{var}(Y)}$

$$\text{var}(X - \alpha Y) = \text{var}(X) - \frac{|\text{cov}(X, Y)|^2}{\text{var}(Y)} = \left(1 - |\rho_{XY}(0)|^2\right) \text{var}(X).$$

- And,

$$\text{var}(\hat{\mu}^{\text{cv}}) = \frac{1 - |\rho_{XY}(0)|^2}{n} \text{var}(X) = \left(1 - |\rho_{XY}(0)|^2\right) \text{var}(\hat{\mu})$$

- So, for this $\alpha = \frac{\text{cov}(X, Y)}{\text{var}(Y)}$

$$\text{var}(X - \alpha Y) = \text{var}(X) - \frac{|\text{cov}(X, Y)|^2}{\text{var}(Y)} = \left(1 - |\rho_{XY}(0)|^2\right) \text{var}(X).$$

- And,

$$\text{var}(\hat{\mu}^{\text{cv}}) = \frac{1 - |\rho_{XY}(0)|^2}{n} \text{var}(X) = \left(1 - |\rho_{XY}(0)|^2\right) \text{var}(\hat{\mu})$$

- How do I apply to spectral estimation?

For timeseries $X = (X_j, j = 1, \dots, N)$,

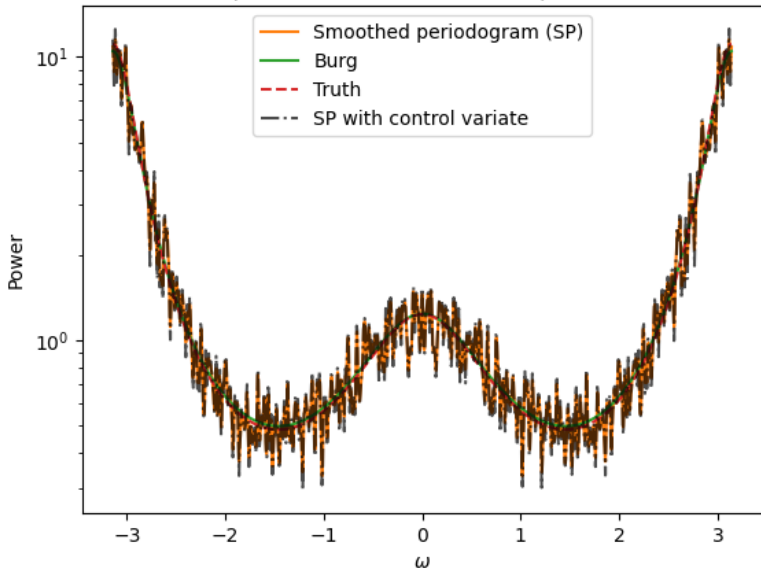
- 1 Divide the full timeseries X into K segments.
- 2 For each segment k , estimate the spectrum $\hat{S}^{(k)}$ and the whitened spectrum $\hat{W}^{(k)}$.
- 3 Take the logarithm $(\log \hat{S}^{(k)})_{k=1}^K$ and $(\log \hat{W}^{(k)})_{k=1}^K$.
- 4 Compute $\alpha = \frac{\text{cov}_k(\log \hat{S}^{(k)}, \log \hat{W}^{(k)})}{\text{var}_k(\log \hat{W}^{(k)})}$, at each frequency.
- 5 For \hat{S} and \hat{W} , the spectrum and whitened spectrum of the full series, put

$$\hat{S}^{\text{CV}} = \exp \left(\log \hat{S} - \alpha \log \hat{W} \right).$$

Example 1: AR(2)



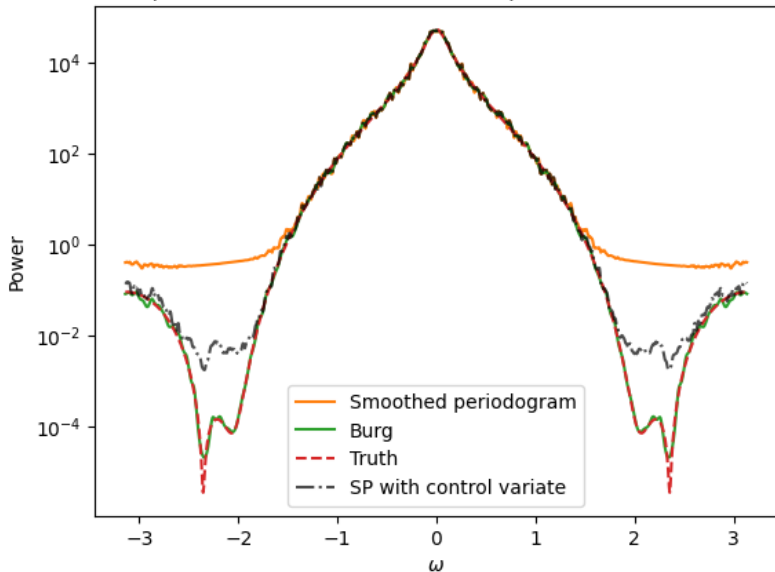
Spectral estimates of AR(2) process



Example 2: AR(10) Signal



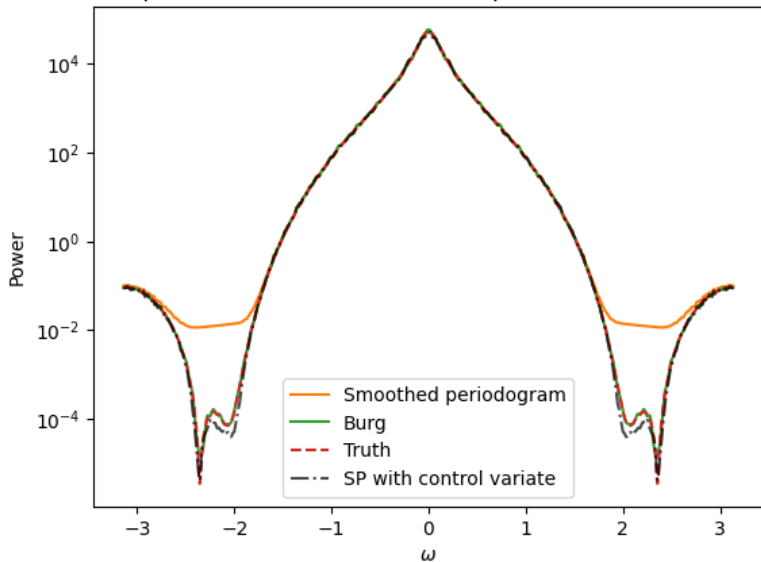
Spectral estimates of ARMA(1,6) process, $N = 10000$



Example 2: AR(10) Signal



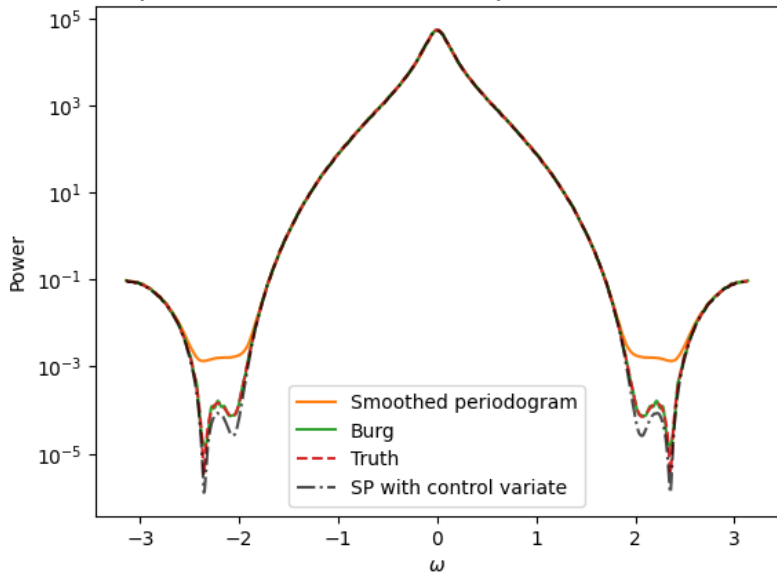
Spectral estimates of ARMA(1,6) process, $N = 100000$



Example 2: AR(10) Signal



Spectral estimates of ARMA(1,6) process, $N = 1000000$



I first saw this while trying to whiten a KSE solution.

The Kuramoto-Sivishinsky equation (KSE) can be written as follows

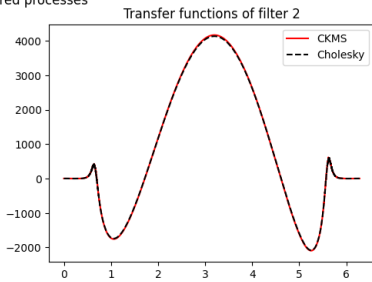
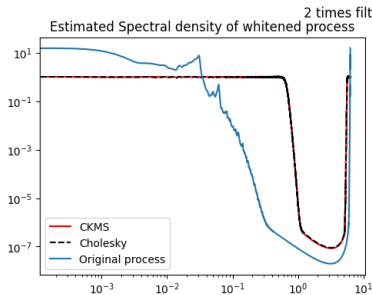
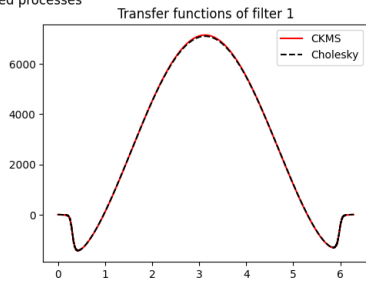
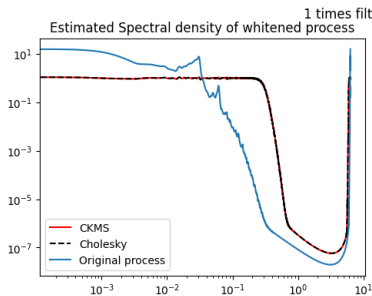
$$u_t + uu_x + u_{xx} + u_{xxxx} = 0$$

with $u(x + L, t) = u(x, t)$ for all $x \in \mathbb{R}$ and $t > 0$. And with $u(x, 0) = g(x)$. Now, we use a fourier series to rewrite the KSE in Fourier space. Doing so gives

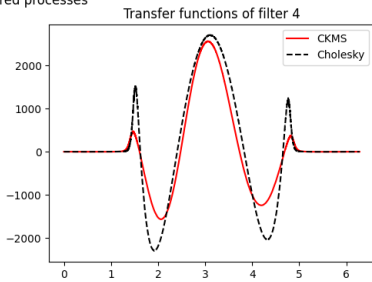
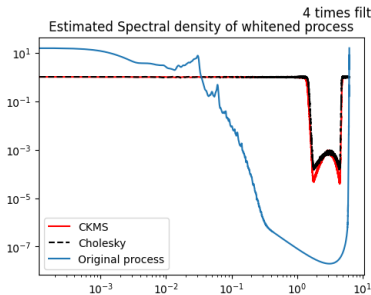
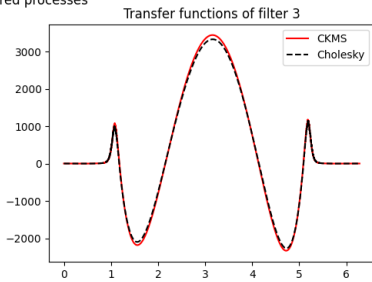
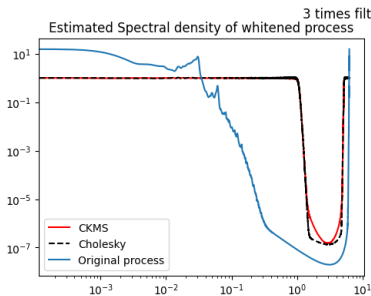
$$\hat{u}_k = (q_k^2 - q_k^4)\hat{u}_k - \frac{iq_k}{2} \sum_{\ell=-\infty}^{\infty} \hat{u}_\ell \hat{u}_{k-\ell} \quad (1)$$

Here, $q_k = \frac{2\pi}{L}k$. Note the trick: $uu_x = \frac{1}{2}(u^2)_x$.

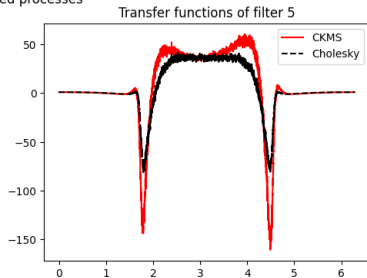
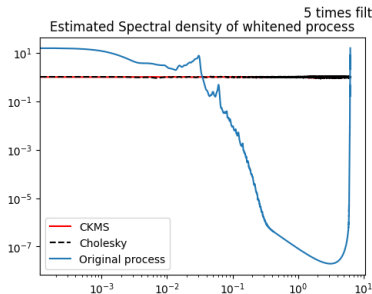
Example 3: KSE



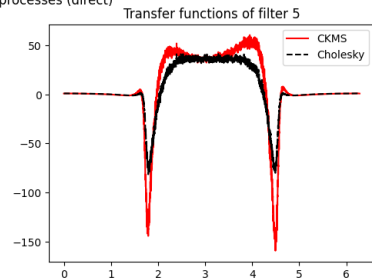
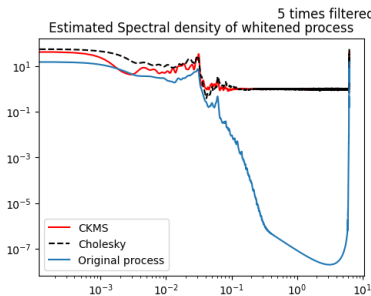
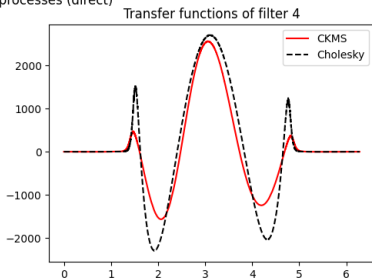
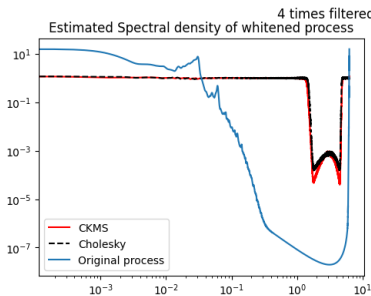
Example 3: KSE



Example 3: KSE



Example 3: KSE



Thank you!



Ali H Sayed and Thomas Kailath.

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Numerical linear algebra with applications, 8(6-7):467–496, 2001.



Thomas Kailath, Ali H Sayed, and Babak Hassibi.

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